

# Probabilistic Graphical Models (5): temporal models

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- Understanding temporal models
- Dynamic Bayesian Network
- Dynamic Markov Random Fields (including CRFs)

# Understanding temporal models

Temporal: **of or relating to time.**

Temporal models are often seen as

- (deterministic) **dynamical systems** (a fixed rule) *e.g.* Kalman filter
- or **random processes** (random walk) *e.g.* HMM, CRFs

# Understanding temporal models

A **dynamical system** uses a **fixed rule** to describe the time dependence of a point in a geometrical space.

- Origin: Newtonian mechanics.
- Often uses **differential equation** *e.g.* a flow  
$$\frac{d}{dt} \mathbf{x}(t) = \mathbf{A} \cdot \mathbf{x}(t),$$
  
or **recurrence relation** *e.g.* Fibonacci numbers  
$$F_n = F_{n-1} + F_{n-2}$$
 with seed values:  $F_0 = 0, F_1 = 1$ .

# Understanding temporal models

**Kalman filter** assumes the true state at time  $t$  is evolved from the state at  $(t - 1)$  according to

$$\mathbf{x}^t = \mathbf{F}^t \mathbf{x}^{t-1} + \mathbf{B}^t \mathbf{u}^t + \mathbf{w}^t,$$

where  $\mathbf{F}^t$  is the state transition model,  $\mathbf{B}^t$  is the control-input model applied to the control vector  $\mathbf{u}^t$ ,  $\mathbf{w}^t$  is the noise from  $\mathcal{N}(0, \mathbf{Q}^t)$ .

Hence it is seen as a dynamical system. Details of Kalman filter and particle filter will be deferred to later part of the talk.

# Understanding temporal models

A **random process** is a collection  $\{X^t : t \in T\}$ , where each  $X^t$  is a random variable.

If  $\Pr(X^{t+1} | X^{0:t}) = \Pr(X^{t+1} | X^t)$ , it's **markovian**.

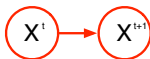
If  $\Pr(X^{t+1} | X^t) = \Pr(X^{t'+1} | X^{t'})$  for all  $t, t'$ , it's **homogenous**.

A **markov chain** is a discrete-time random process with markovian and homogenous assumptions.

A chain of length  $N + 1$

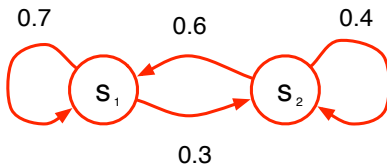


Essentially modeling (due to homogenous assumption)



# Understanding temporal models

**Transition graph** (assuming  $X \in \{s_1, s_2\}$ ) of the previous markov chain is:

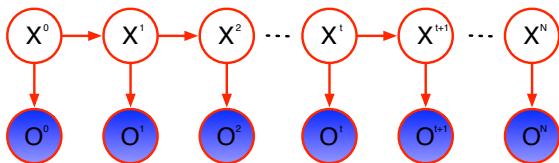


Transition graph  $\neq$  graphical model

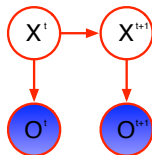
# Understanding temporal models

HMM (with hidden state  $X^t$  and observation  $O^t$ ):

A HMM chain of length  $N + 1$



Essentially modeling (due to homogenous assumption)



This is called **2-time-slice Bayesian network (2-TBN)**.



# Dynamic Bayesian Network

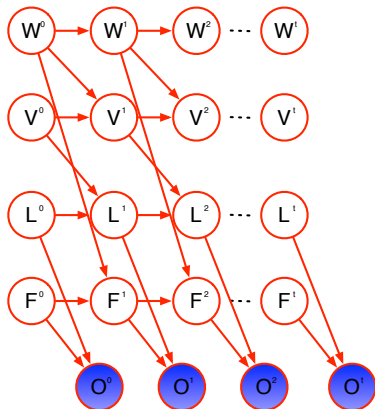
A **Dynamic Bayesian Network** (DBN) is a pair  $(\mathcal{B}_0, \mathcal{B}_{\rightarrow})$ , where  $\mathcal{B}_0$  is a Bayesian network over  $\mathbf{Z}^0$  representing the initial distribution over states, and  $\mathcal{B}_{\rightarrow}$  is a 2-TBN for the process representing the transition between states over  $\mathbf{Z}^{t:t+1}$ . For any  $t \geq 0$ , the distribution over  $\mathbf{Z}^{0:t}$  is a **unrolled Bayesian network**. See an example.

# Dynamic Bayesian Network

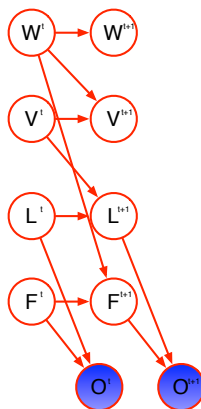
A moving car tries to track its current location using the data obtained from a possibly faulty sensor with states set {Weather, Velocity, Location, Failure, Observation}

$$\mathbf{Z}^t = (W^t, V^t, L^t, F^t, O^t)$$

An unrolled DBN (over  $\mathbf{Z}^{0:t}$ )



2-TBN ( $\mathcal{B}_{\rightarrow}$ )

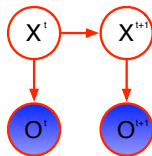


# Dynamic Bayesian Network

Given training data, the **parameters** of the DBN can be learnt via techniques in tutorial (3). Once parameters are learnt, the **prediction** can be done via techniques in tutorial (2).

# Dynamic Bayesian Network

Kalman filter as a HMM.



Markov assumption:  $p(\mathbf{x}^{t+1} | \mathbf{x}^0, \dots, \mathbf{x}^t) = p(\mathbf{x}^{t+1} | \mathbf{x}^t)$  and  $p(\mathbf{o}^{t+1} | \mathbf{x}^0, \dots, \mathbf{x}^{t+1}) = p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1})$

$$\text{MAP: } p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t+1}) = \frac{p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1})p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t})}{p(\mathbf{o}^{t+1} | \mathbf{o}^{0:t})},$$

$$\text{where } p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t}) = \int p(\mathbf{x}^{t+1}, \mathbf{x}^t | \mathbf{o}^{0:t}) d\mathbf{x}^t = \int p(\mathbf{x}^{t+1} | \mathbf{x}^t) p(\mathbf{x}^t | \mathbf{o}^{0:t}) d\mathbf{x}^t,$$

$$\text{and } p(\mathbf{o}^{t+1} | \mathbf{o}^{0:t}) = \int p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1}) p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t}) d\mathbf{x}^{t+1} \quad \text{constant w.r.t. } \mathbf{x}$$

The remaining terms are ready to compute:

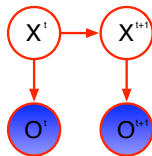
$$p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1}) = \mathcal{N}(\mathbf{H}^{t+1} \mathbf{x}^{t+1}, \mathbf{R}^{t+1})$$

$$p(\mathbf{x}^{t+1} | \mathbf{x}^t) = \mathcal{N}(\mathbf{F}^{t+1} \mathbf{x}^t, \mathbf{Q}^{t+1})$$

$$p(\mathbf{x}^t | \mathbf{o}^{0:t}) = \mathcal{N}(\hat{\mathbf{X}}^t, \mathbf{P}^t)$$

# Dynamic Bayesian Network

## Particle filter as a HMM.



Markov assumption:  $p(\mathbf{x}^{t+1} | \mathbf{x}^0, \dots, \mathbf{x}^t) = p(\mathbf{x}^{t+1} | \mathbf{x}^t)$  and  $p(\mathbf{o}^{t+1} | \mathbf{x}^0, \dots, \mathbf{x}^{t+1}) = p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1})$

$$\text{MAP: } p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t+1}) = \frac{p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1})p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t})}{p(\mathbf{o}^{t+1} | \mathbf{o}^{0:t})},$$

where  $p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t}) = \int p(\mathbf{x}^{t+1}, \mathbf{x}^t | \mathbf{o}^{0:t}) d\mathbf{x}^t = \int p(\mathbf{x}^{t+1} | \mathbf{x}^t) p(\mathbf{x}^t | \mathbf{o}^{0:t}) d\mathbf{x}^t = \mathbb{E}_{\mathbf{x}^t \sim p(\mathbf{x}^t | \mathbf{o}^{0:t})} [p(\mathbf{x}^{t+1} | \mathbf{x}^t)],$

$$\text{and } p(\mathbf{o}^{t+1} | \mathbf{o}^{0:t}) = \int p(\mathbf{o}^{t+1} | \mathbf{x}^{t+1}) p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t}) d\mathbf{x}^t$$

When  $p(\mathbf{x}^t | \mathbf{o}^{0:t})$  is not gaussian, one can use monte carlo ( $N$  samples  $\{\mathbf{x}_i^t\}_{i=1}^N$ ) to approximate the expectation

$$\mathbb{E}_{\mathbf{x}^t \sim p(\mathbf{x}^t | \mathbf{o}^{0:t})} [f(\mathbf{x}^t)] \approx \frac{1}{N} \left[ \sum_{i=1}^N f(\mathbf{x}_i^t) \right],$$

where  $f(\mathbf{x}^t) = p(\mathbf{x}^{t+1} | \mathbf{x}^t).$

# Dynamic Bayesian Network

Particle filter uses **Sequential Importance Resampling** (SIR)

Step 1: **sampling and computing weights.**

Sample  $\mathbf{x}_i^t \sim q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t}) \approx p(\mathbf{x}^t | \mathbf{x}^{t-1})$ ,

$$w_i^t = w_i^{t-1} \cdot \frac{p(\mathbf{o}^t | \mathbf{x}_i^t) p(\mathbf{x}_i^t | \mathbf{x}_i^{t-1})}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})}.$$

Step 2: **resampling.** If not enough particles  $\{\mathbf{x}_i^t\}_i$ , resample  $\{\mathbf{y}_i^t\}_i$  from the current particle set  $\{\mathbf{x}_i^t\}_i$  with probability proportional to  $\{w_i\}_i$ . Replace the current particle set with the new one  $\{\mathbf{y}_i^t\}_i$ .

$$\begin{aligned} p(\mathbf{x}^{t+1} | \mathbf{o}^{0:t}) &= \mathbb{E}_{\mathbf{x}^t \sim p(\mathbf{x}^t | \mathbf{o}^{0:t})} [p(\mathbf{x}^{t+1} | \mathbf{x}^t)], \\ &\approx \frac{1}{\sum_{i=1}^N w_i^t} \left[ \sum_{i=1}^N w_i^t p(\mathbf{x}^{t+1} | \mathbf{x}_i^t) \right] \approx \frac{1}{N} \left[ \sum_{i=1}^N p(\mathbf{x}^{t+1} | \mathbf{y}_i^t) \right] \end{aligned}$$

The way that  $w_i^t$  is generated may look **strange** at first sight.  
Let's crack it!

# Dynamic Bayesian Network

In **importance sampling**, one wants to sample (but hard to)  $\mathbf{x} \sim p(\mathbf{x})$ , and it's easy to sample  $\mathbf{x} \sim q(\mathbf{x})$ .

$$\mathbb{E}_p[f(\mathbf{x})] = \mathbb{E}_q\left[\frac{p(\mathbf{x})}{q(\mathbf{x})} \cdot f(\mathbf{x})\right] \approx \frac{\sum_{i=1}^N w_i \cdot f(\mathbf{x}_i)}{\sum_{i=1}^N w_i}$$

The weight

$$w_i = \frac{p(\mathbf{x}_i)}{q(\mathbf{x}_i)}.$$

In particle filter, at time  $t$  and  $t - 1$ ,

$$w^t = \frac{p(\mathbf{x}^{0:t} | \mathbf{o}^{0:t})}{q(\mathbf{x}^{0:t} | \mathbf{o}^{0:t})},$$

$$w^{t-1} = \frac{p(\mathbf{x}^{0:t-1} | \mathbf{o}^{0:t-1})}{q(\mathbf{x}^{0:t-1} | \mathbf{o}^{0:t-1})}.$$

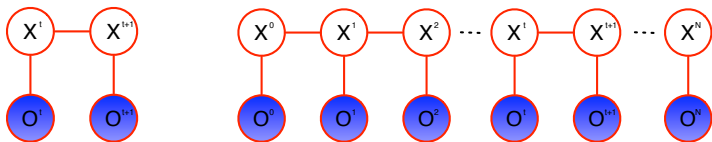
# Dynamic Bayesian Network

$$\begin{aligned}\frac{w^t}{w^{t-1}} &= \frac{p(\mathbf{x}^{0:t} | \mathbf{o}^{0:t}) q(\mathbf{x}^{0:t-1} | \mathbf{o}^{0:t-1})}{q(\mathbf{x}^{0:t} | \mathbf{o}^{0:t}) p(\mathbf{x}^{0:t-1} | \mathbf{o}^{0:t-1})} \\ &= \frac{p(\mathbf{x}^t, \mathbf{o}^t, \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t-1})}{p(\mathbf{x}^{0:t-1}, \mathbf{o}^{0:t-1})} \cdot \frac{p(\mathbf{o}^{0:t-1})}{p(\mathbf{o}^{0:t})} \cdot \frac{1}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})} \\ &= p(\mathbf{x}^t, \mathbf{o}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t-1}) \cdot \frac{p(\mathbf{o}^{0:t-1})}{p(\mathbf{o}^{0:t})} \cdot \frac{1}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})} \\ &= p(\mathbf{x}^t, \mathbf{o}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t-1}) \cdot \text{const} \cdot \frac{1}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})} \\ &= p(\mathbf{x}^t, \mathbf{o}^t | \mathbf{x}^{t-1}, \mathbf{o}^{t-1}) \cdot \text{const} \cdot \frac{1}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})} \\ &= p(\mathbf{x}^t, \mathbf{o}^t | \mathbf{x}^{t-1}) \cdot \text{const} \cdot \frac{1}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})} \\ \therefore w^t &= w^{t-1} \cdot \frac{p(\mathbf{o}^t | \mathbf{x}^t) p(\mathbf{x}^t | \mathbf{x}^{t-1})}{q(\mathbf{x}^t | \mathbf{x}^{0:t-1}, \mathbf{o}^{0:t})} \cdot \text{const}.\end{aligned}$$



# Dynamic Markov Random Fields

CRFs:



$$\mathbf{P}(\mathbf{y} | \mathbf{x}; \mathbf{w}) = \frac{\exp(\langle \mathbf{w}, \Phi(\mathbf{x}, \mathbf{y}) \rangle)}{Z(\mathbf{w} | \mathbf{x})}, \quad (1)$$

where

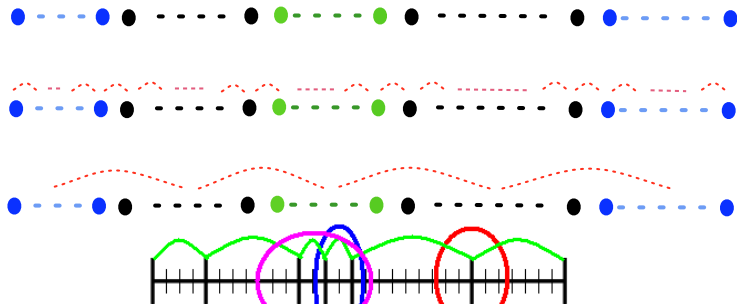
$$Z(\mathbf{w} | \mathbf{x}) = \sum_{\mathbf{y}' \in \mathcal{Y}} \exp(\langle \mathbf{w}, \Phi(\mathbf{x}, \mathbf{y}') \rangle), \quad (2)$$

and

$$\Phi(\mathbf{x}, \mathbf{y}) = \sum_{i \in \mathcal{V}} \Phi_1(\mathbf{x}, \mathbf{y}^{(i)}) + \sum_{(ij) \in \mathcal{E}} \Phi_2(\mathbf{x}, \mathbf{y}^{(ij)}). \quad (3)$$

# Dynamic Markov Random Fields

Semi-Markov Model: Using (??) with



$$\Phi(\mathbf{x}, \mathbf{y}) = \left( \sum_{i=0}^{l-1} \Phi_1(\mathbf{x}, n_i, c_i), \sum_{i=0}^{l-1} \Phi_2(\mathbf{x}, n_i, n_{i+1}, c_i), \sum_{i=0}^{l-1} \Phi_3(\mathbf{x}, n_i, n_{i+1}, c_i, c_{i+1}) \right).$$

$\phi_3$      $\phi_2$      $\phi_1$   
—    —    —